

De De De De De

Modeling Multi-country Mortality Dependence by a Vine Copula

Masafumi Suzuki

FIAJ, CERA The Prudential Gibraltar Financial Life Insurance Co., Ltd.

```
May 11<sup>th</sup> – May 15<sup>th</sup> 2020
```

About the speaker





Masafumi Suzuki

- Manager, Actuarial Team
- Fellow of The Institute of Actuaries of Japan(FIAJ), CERA
- Member of ASTIN-related Study Group of The Institute of Actuaries of Japan.



The Prudential Gibraltar Financial Life Insurance Co., Ltd.

- Japanese subsidiary of Prudential Financial, Inc.
- Bancassurance Channel



Introduction

- Multi-country mortality dependence attracts the attention of insurers operating life insurance or annuity business in multiple countries.
- When implementing a sophisticated enterprise risk management (ERM) program, it is crucial to model the structure of such dependence accurately.
- Elliptic and Archimedean copulas are often used for risk aggregation in advanced ERM.
- However, these well-known copulas cannot always flexibly capture complex tail dependence, especially under certain stressed situations.
- This study proposes modeling multi-country mortality dependence by a vine copula, which provides greater flexibility and efficiently characterizes the dependence structure.

SECTIONS VIRTUAL COLLOQUIUM 2020

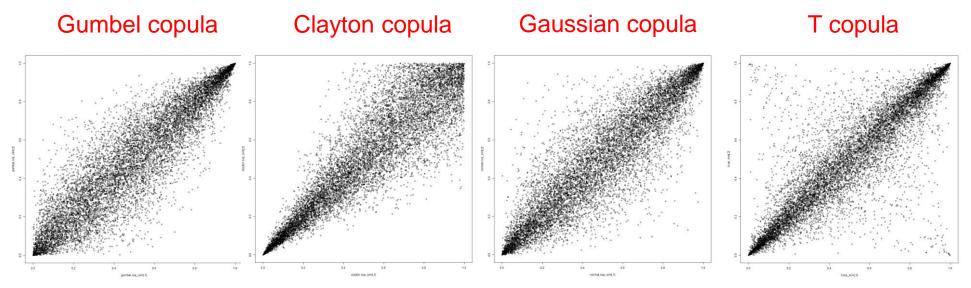
What is a copula ?

■ A copula is the joint distribution of distribution functions of random valuables.

 $F(x_1, x_2) = C(F_1(x_1), F_2(x_2)).$

A copula is able to flexibly express the dependence structure between bivariate random valuables, especially the tail dependence.

■ Follows are the plots of well-known bivariate copulas(Kendall's tau =0.75).





Limitation of well-known copulas

- Archimedean copulas(such as Gumbel copula and Clayton copula) have only one parameter, so these copulas cannot express the different dependence between each pair of random variables.
- Gaussian copula does not have tail dependence and underestimates the risk.
- T copula is better, but not flexible because t copula expresses the tail dependence by only one parameter (the degree of freedom).

Copulas	# parameters	Lower tail dependence	Upper tail dependence
Gumbel copula	1	0	$2 - 2^{1/\alpha}$
Clayton copula	1	$2^{-1/\alpha}$	0
Gaussian copula	n(n-1)/2	0	0
T copula	n(n-1)/2+1	$2t_{\nu+1}\left(-\sqrt{\frac{(\nu+1)(1-\rho)}{1+\rho}}\right)$	$2t_{\nu+1}\left(-\sqrt{\frac{(\nu+1)(1-\rho)}{1+\rho}}\right)$

n : the number of variables v : the degree of freedom ρ : correlation coefficient



What is a vine copula ?

In applying Sklar's theorem proposed by Sklar(1959), one can construct the joint probability density function of random variables X_1 and X_2 as follows:

$$f(x_1, x_2) = c_{1,2} \big(F_1(x_1), F_2(x_2) \big) f_1(x_1) f_2(x_2).$$
(1)

We can also express the joint probability density function using conditional probability as follows:

$$f(x_1, x_2) = f_{1|2}(x_1|x_2) f_2(x_2).$$
⁽²⁾

Combining equation (1) and (2), $f_{1|2}(x_1|x_2)$ is shown as follows:

$$f_{1|2}(x_1|x_2) = c_{1,2}(F_1(x_1), F_2(x_2))f_1(x_1).$$
(3)



What is a vine copula ?

We can express the joint probability density function of random valuables (X_1 , X_2 and , X_3) using the conditional probability density function as follows:

$$f(x_1, x_2, x_3) = f_3(x_3) f_{2|3}(x_2|x_3) f_{1|2,3}(x_1|x_2, x_3).$$
(1)

Following the last page, $f_{2|3}(x_2|x_3)$ can be written as follows:

$$\boldsymbol{f_{2|3}(x_2|x_3)} = c_{2,3}(F_2(x_2), F_3(x_3))f_2(x_2).$$
(2)

 $\begin{aligned} f_{1|2,3}(x_1|x_2, x_3) & \text{ can be obtained by a conditional bivariate copula density function} \\ f_{1|2,3}(x_1|x_2, x_3) &= c_{1,3|2} \left(F_{1|2}(x_1|x_2), F_{3|2}(x_3|x_2) \right) f_1(x_1|x_2) \\ &= c_{1,3|2} \left(F_{1|2}(x_1|x_2), F_{3|2}(x_3|x_2) \right) c_{1,2} \left(F_{1}(x_1), F_{2}(x_2) \right) f_1(x_1). \end{aligned}$ (3)

Combining (1),(2), and (3), we obtain

$$f(x_1, x_2, x_3) = c_{1,3|2} \left(F_{1|2}(x_1|x_2), F_{3|2}(x_3|x_2) \right) c_{1,2} \left(F_1(x_1), F_2(x_2) \right)$$
(4)

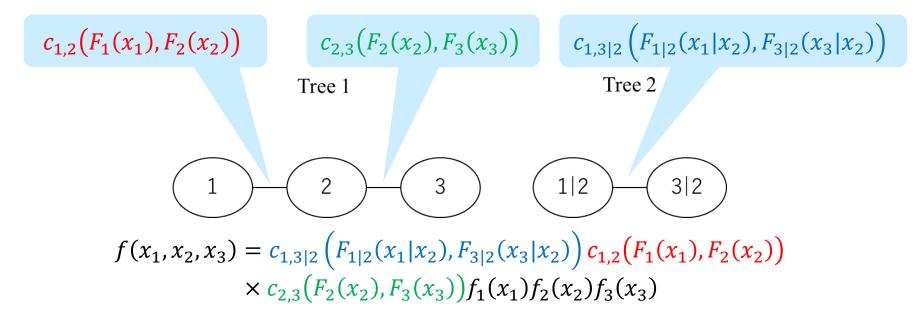
$$\times c_{2,3} \left(F_2(x_2), F_3(x_3) \right) f_1(x_1) f_2(x_2) f_3(x_3).$$

This feature can be easily generalized to n dimensions.



What is a vine copula ?

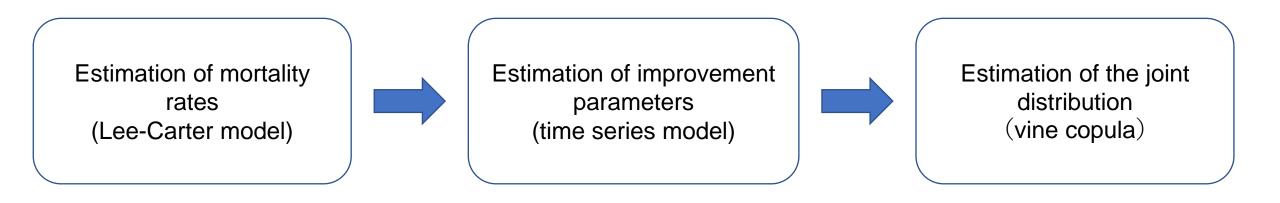
- Such dependence structure can be graphically expressed by vines, the concept of graph theory.
- A vine copula expresses the node as a random variable and the edge as a conditional bivariate copula.
- The application for market risk and credit risk modeling has been studied in recent years.





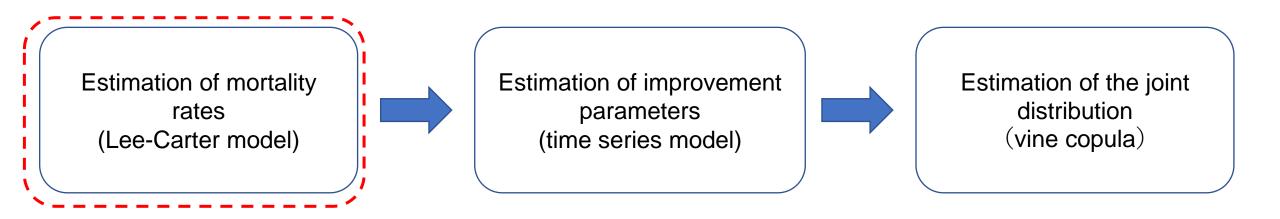
The overview of this research

- Express the mortality rate of 13 countries(12 European countries and Australia) by a Lee-Carter model.
- Model the joint distribution of mortality improvement parameters of each country by a vine copula.
- Compare our model to the models by benchmark copulas.





Estimation of mortality rates





Data & Methodology

Data source	:	Human Mortality Database
Countries	:	12 European countries + Australia(AUS)
Age	:	From 65 to 90 inclusive
Sex	:	Male
Year	:	$1921 \sim 2014$

- We model the marginal dynamic of the mortality rates for each country based on a Lee-Carter model.
- In the modeling of this study, we use R package
 - "StMoMo" and "VineCopula".





Lee-Carter model

Mortality sensitivity by κ_t

$$log(q_{x,t}) = \alpha_x + \beta_x \cdot \kappa_t + e_{x,t}$$
Logarithm of Base q_x Improvement parameter by year

- $q_{x,t}$: base mortality rates at age x and time t
- α_x : average profile of mortality
- κ_t : mortality changes over time

 β_x : how much each age group mortality changes when κ_t^J changes $e_{x,t}$: error term which reflects the effects not captured by the model



Estimation of morality rates of 13 countries

$$log(q_{x,t}^{1}) = \alpha_{x}^{1} + \beta_{x}^{1}\kappa_{t}^{1} + e_{x,t}^{1}$$
 Australia

$$log(q_{x,t}^{2}) = \alpha_{x}^{2} + \beta_{x}^{2}\kappa_{t}^{2} + e_{x,t}^{2}$$
 Belgium

$$log(q_{x,t}^{3}) = \alpha_{x}^{3} + \beta_{x}^{3}\kappa_{t}^{3} + e_{x,t}^{3}$$
 Nederland

$$log(q_{x,t}^{4}) = \alpha_{x}^{4} + \beta_{x}^{4}\kappa_{t}^{4} + e_{x,t}^{4}$$
 England & Wales

$$log(q_{x,t}^{5}) = \alpha_{x}^{5} + \beta_{x}^{5}\kappa_{t}^{5} + e_{x,t}^{5}$$
 Denmark

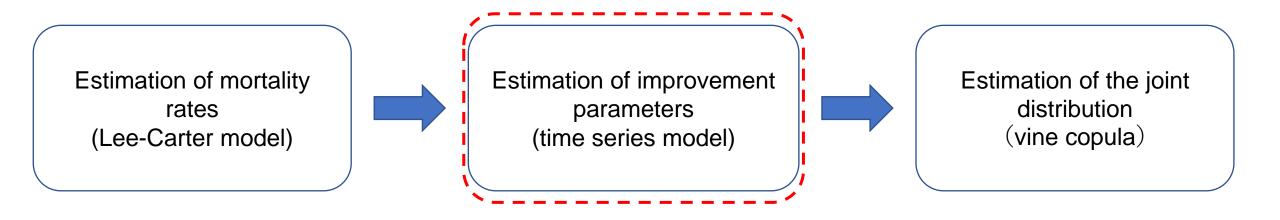
$$\vdots$$

$$log(q_{x,t}^{13}) = \alpha_{x}^{13} + \beta_{x}^{13}\kappa_{t}^{13} + e_{x,t}^{13}$$
 Iceland

We fit the Lee-carter model to the mortality data of each country.



Estimation of improvement parameters





Estimation of future mortality rates

$$log(q_{x,t}^{1}) = \alpha_{x}^{1} + \beta_{x}^{1} \kappa_{t}^{1} + e_{x,t}^{1}$$
Australia

$$log(q_{x,t}^{2}) = \alpha_{x}^{2} + \beta_{x}^{2} \kappa_{t}^{2} + e_{x,t}^{2}$$
Belgium

$$log(q_{x,t}^{3}) = \alpha_{x}^{3} + \beta_{x}^{3} \kappa_{t}^{3} + e_{x,t}^{3}$$
Nederland

$$log(q_{x,t}^{4}) = \alpha_{x}^{4} + \beta_{x}^{4} \kappa_{t}^{4} + e_{x,t}^{4}$$
England & Wales

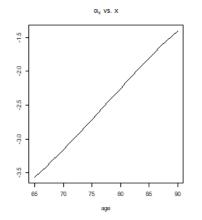
$$log(q_{x,t}^{5}) = \alpha_{x}^{5} + \beta_{x}^{5} \kappa_{t}^{5} + e_{x,t}^{5}$$
Denmark

$$\vdots$$

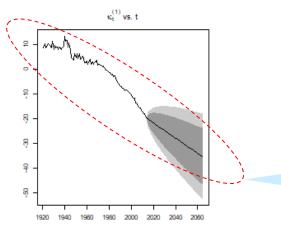
$$log(q_{x,t}^{13}) = \alpha_{x}^{13} + \beta_{x}^{13} \kappa_{t}^{13} + e_{x,t}^{13}$$
Iceland

We assume α_x^j and β_x^j are constant over the future period, and model κ_t^j as a time series model.

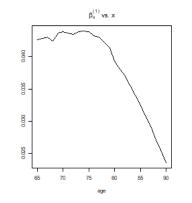
The result of estimation (ex. France)



 α_x : Logarithm of base q_x



 κ_t : Improvement parameter by year



 β_x : Mortality sensitivity by κ_t

$$log(q_{x,t}) = \alpha_x + \beta_x \cdot \kappa_t + e_{x,t}$$

The movement of κ_t looks straight \Rightarrow Random walk with a drift term $\kappa_t = \kappa_{t-1} + \mu + \varepsilon_t$

ACTUAIRES

Modeling the movement of the improvement parameter

$$\kappa_{t}^{1} = \kappa_{t-1}^{1} + \mu^{1} + \varepsilon_{t}^{1}, \ \varepsilon_{t}^{1} \sim N(0, \sigma_{1}^{2})$$

$$\kappa_{t}^{2} = \kappa_{t-1}^{2} + \mu^{2} + \varepsilon_{t}^{2}, \ \varepsilon_{t}^{2} \sim N(0, \sigma_{2}^{2})$$

$$\kappa_{t}^{3} = \kappa_{t-1}^{3} + \mu^{3} + \varepsilon_{t}^{3}, \ \varepsilon_{t}^{3} \sim N(0, \sigma_{3}^{2})$$

$$\kappa_{t}^{4} = \kappa_{t-1}^{4} + \mu^{4} + \varepsilon_{t}^{4}, \ \varepsilon_{t}^{4} \sim N(0, \sigma_{4}^{2})$$

$$\kappa_{t}^{5} = \kappa_{t-1}^{5} + \mu^{5} + \varepsilon_{t}^{5}, \ \varepsilon_{t}^{5} \sim N(0, \sigma_{5}^{2})$$

$$\vdots$$

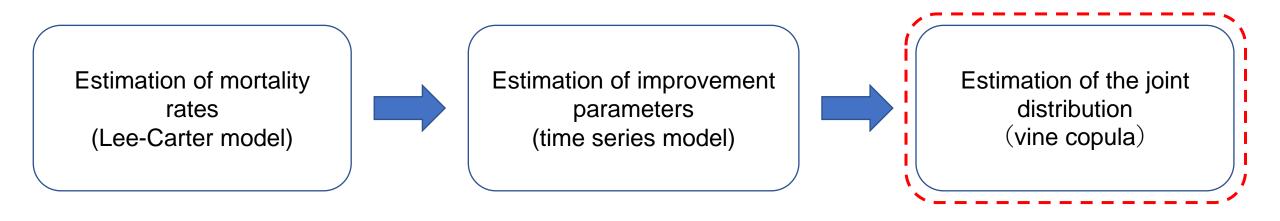
$$\kappa_{t}^{13} = \kappa_{t-1}^{13} + \mu^{13} + \varepsilon_{t}^{13}, \ \varepsilon_{t}^{13} \sim N(0, \sigma_{13}^{2})$$

Each error term follows a normal distribution.

In this study, we model the joint distribution of ε_t^i by a vine copula.

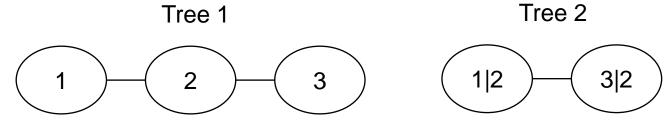


Estimation of the joint distribution

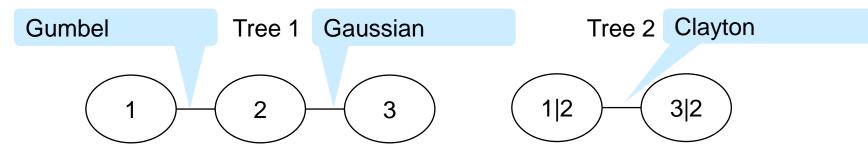




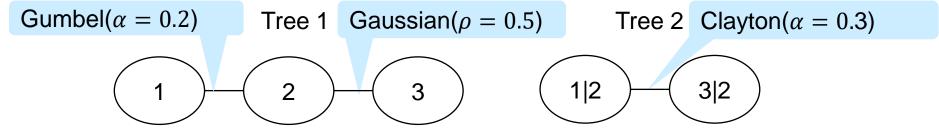
1. Determine the tree structure.



2. Determine the copula families applied to each edge of the vine.



3. Estimate the parameters of each copula.





Modeling by a vine copula

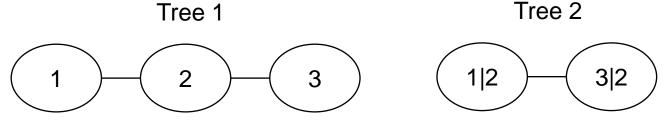
- Although a vine copula has flexibility, the number of tree structures is too huge to choose.
- The following table illustrates the super-exponential growth of the number of structures for dimensions n = 2, ..., 10.
- An automatic algorithm is necessary for modeling by a vine copula.

Dimension (n)	# tree structures	Tree × copula*
2	1	7
3	3	1,029
4	24	2,823,576
5	480	1.3559 e+11
6	23,040	1.0938 e+17
7	2,580,480	1.4413 e+24
8	660,602,880	3.0387 e+32
9	3.8051e+11	1.0090 e+42
10	4.8705e+14	5.2118 e+52

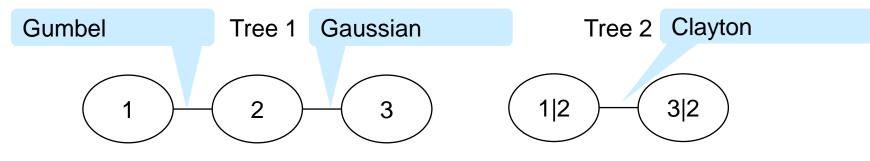
*When there are 7 options for copula.

Sequential method proposed by Dißmann et al. (2013)

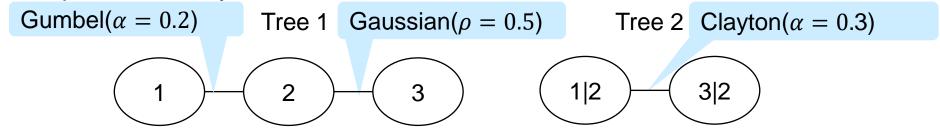
1. Choose the tree structure that maximizes the absolute value of each pair's Kendall's tau.



2. Select the bivariate copula for each edge of the tree based on AIC.



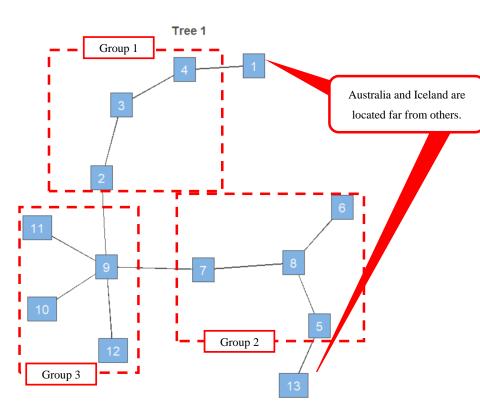
3. Estimate the parameters by maximum likelihood estimation.

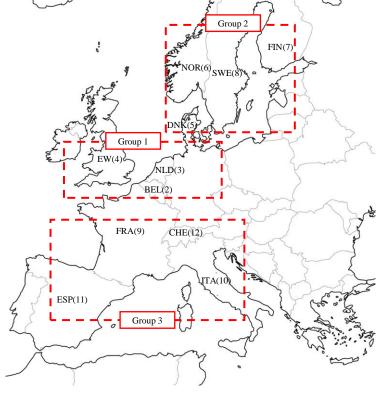




Result and consideration

Dependence structure by the vine copula





1 <-> AUS, 2 < -> BEL, 3 <-> NLD, 4 < -> EW, 5 < -> DNK, 6 < -> NOR, 7 < -> FIN, 8 <-> SWE, 9 < -> FRA, 10 < -> ITA, 11 < -> ESP, 12 <-> CHE, 13 < -> ICE

By using the sequential method, we obtain a dependence structure resembling the actual geographical relationships, which is intuitively understandable.



Copula families and parameters

Edge	Copula	Parameter1	Parameter2*1	Kendall's τ	λ _U *2	λ_L^*2
4-1	Т	0.10	3.83	0.07	0.11	0.11
3-4	Т	0.75	6.85	0.50	0.28	0.28
2-3	Т	0.80	5.13	0.59	0.44	0.44
9-11	Frank	4.47	I	0.42	-	-
9-2	Frank	8.74	I	0.63	-	-
5-13	Frank	1.68	-	0.18	-	-
8-6	Gumbel	1.59	I	0.37	0.45	-
8-5	Gaussian	0.65	I	0.45	-	-
7-8	Frank	4.98	-	0.46	-	-
9-7	Т	0.64	30.00	0.44	0.01	0.01
9-10	Serval	2 2 2		0.57		0.65
9-10	Gumbel	2.32	-	0.57	-	0.65
12-9	Frank	8.13	-	0.61	-	

*1 The parameter2 of t copula is a degree of freedom.

*2 λ_U and λ_L denote upper and lower tail-dependence coefficient respectively.

The different copula is applied to each pair of variables.

The vine copula model can capture complex dependence, especially the tail dependence.²⁴

Comparison with benchmark copulas

The performance of the vine copula is the best in all criteria.

- The Frank Copula and the Clayton Copula have only one parameter, and cannot express the difference by the pairs of variables.
- The Gaussian copula and the t copula can set different correlation coefficients for each pair of variables, but the dependence is symmetrical and linear.

	Log-likelihood ^{*1}	AIC ^{*2}	BIC ^{*2}
Vine copula	415.4	-702.9	-540.8
Gaussian copula	313.3	-476.4	-273.1
T copula	342.4	-526.8	-326.7
Frank copula	112.2	-222.5	-219.9
Clayton copula	68.6	-135.1	-132.6

*1 Larger is better.

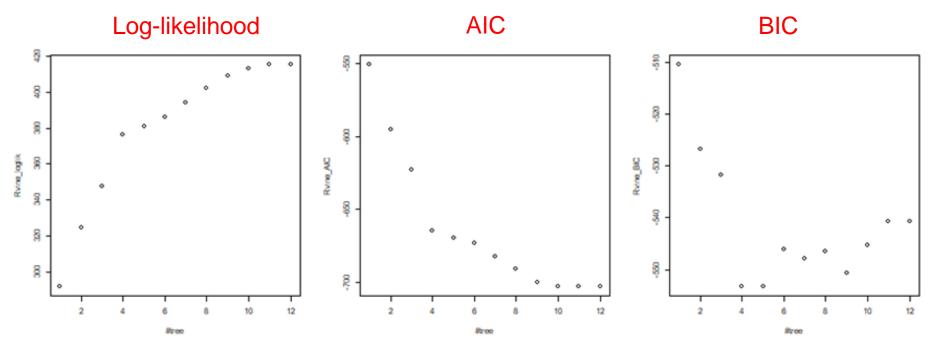
*2 Smaller is better.

Analysis of the contribution by copulas in tree 2 and above

In order to investigate how the copulas associated with the higher trees contribute, We set vine copula models consisted of the low-dimensional trees and plot the

log-likelihood, AIC, and BIC.

BIC is the best when we assume independence for the tree 5 and above.





Simulation



Simulation

We defined a survival index by the equation below,

$$_{90}S_{65} = \sum_{j=1}^{n} w_j \prod_{x=65}^{90} \left(1 - q_x^j\right)$$

- Following the capital regulations, such as ICS and Solvency II, we evaluate 99.5% Value at Risk (VaR) over the one year time horizon.
- Estimate the movement of ${}_{90}S_{65}$ after 1 Yr (2,000 times) and evaluate the difference by copula models(vine, Gaussian, and t) in the stress situation.
- We conduct the simulation of the following two cases.
- Case1: All exposures are evenly diversified in all countries.
- Case2: All exposures are concentrated in Group 1 (Belgium(BEL), Netherlands (NLD), and England & Wales(EW)).

Case 1 All exposures are evenly diversified in all countries

■ No significant difference in the 95% ile and 5% ile. (red square)

- The 99.5% ile value of the t copula is the largest, the vine copula is middle, and the Gaussian copula is the smallest.(blue square)
- The Gaussian copula is minimum due to its no tail dependence.
- The vine copula consists of a pair of copulas with and without tail dependence, so the tail risk is larger than a Gaussian copula.
- The t copula shows the highest tail dependence because the dependence of all countries increases in the tail environment.

t copula > vine copula> Gaussian copula			No significant difference			
		99.5%ile	95%ile	50%ile	5%ile	0.5%ile
	Vine copula	1.114	1.060	1.000	0.941	0.890
	Gaussian copula	1.111	1.060	1.000	0.944	0.889
	T copula	1.122	1.058	1.000	0.944	0.893

Case 2 All exposures are concentrated in Group 1

- Similar to case 1, no significant difference in the 95%ile and 5%ile.(red square)
- However, different from case 1, the 99.5% ile value estimated by the vine copula is the largest. (blue square)
- All pair of the countries in Group 1 are modeled by the t copula with stronger tail dependence than the t copula model.
- We underestimate the dependence of the tail of concentrated exposures by Gaussian copula or t copula.

vine copula > t copula > Gaussian copula			No signific	ant differenc	e	
		99.5%ile	95%ile	50%ile	5%ile	0.5%ile
	Vine copula	1.233	1.102	1.000	0.896	0.808
	Gaussian copula	1.193	1.105	1.000	0.900	0.812
	T copula	1.219	1.103	1.000	0.900	0.806



Summary

- We demonstrate the usefulness of a vine copula using actual data by following steps,
 - Use a Lee-Carter model to estimate the mortality rates of 13 countries (12 European countries and Australia).
 - Model the dependence among the time-varying mortality improvement parameters of each country by the vine copula.
 - Demonstrate that the vine copula is superior on some measures (Log-likelihood, AIC, and BIC) to other benchmark copulas.
- We obtain a dependence structure resembling the actual geographical relationships, which is intuitively understandable.
- This study reveals that we underestimate the dependence of the tail risk of concentrated exposures in multiple countries by Gaussian copula or t copula.

Thank you for your attention



Contact details :

Masafumi Suzuki The Prudential Gibraltar Financial Life Insurance Co., Ltd.

masafumi_suzuki@pgf-life.co.jp

<Acknowledgements>
This study has been assisted by ASTIN-related Study Group of The Institute of Actuaries of Japan. I would like to thank its members for their helpful comments and suggestions.



Disclaimer:

The views or opinions expressed in this presentation are those of the authors and do not necessarily reflect official policies or positions of the Institut des Actuaires (IA), the International Actuarial Association (IAA) and its Sections.

While every effort has been made to ensure the accuracy and completeness of the material, the IA, IAA and authors give no warranty in that regard and reject any responsibility or liability for any loss or damage incurred through the use of, or reliance upon, the information contained therein. Reproduction and translations are permitted with mention of the source.

Permission is granted to make brief excerpts of the presentation for a published review. Permission is also granted to make limited numbers of copies of items in this presentation for personal, internal, classroom or other instructional use, on condition that the foregoing copyright notice is used so as to give reasonable notice of the author, the IA and the IAA's copyrights. This consent for free limited copying without prior consent of the author, IA or the IAA does not extend to making copies for general distribution, for advertising or promotional purposes, for inclusion in new collective works or for resale.

References(1/2)

[1] Andres M. Villegas, Pietro Millossovich, Vladimir K. Kaishev (2017) : "StMoMo: An R Package for Stochastic Mortality Modeling", (https://cran.r-project.org/web/packages/StMoMo/).

[2] Chou-Wen Wang, Sharon S.Yangb and Hong-Chih Huang (2015) : "Modeling multi-country mortality dependence and its application in pricing survivor index swaps—A dynamic copula approach" Insurance: Mathematics and Economics 63 (2015) 30–39

[3] DIßMANN, J., BRECHMANN, E.C., CZADO, C. and KUROWICKA, D. (2013) Selecting and estimating regular vine copulae and application to financial returns. Computational Statistics & Data Analysis, 59, 52–69.

[4] Hua Chen, Richard MacMinn and Tao Sun (2013) :"Multi-population mortality models: A factor copula approach", Fox School of Business Research Paper, No. 15-052

[5] Helena Chuliá, Montserrat Guillén and Jorge Uribe(2016) : "Modeling Longevity Risk with Generalized Dynamic Factor Models and Vine-Copulae", ASTIN Bulletin: The Journal of the International Actuarial Association, 2016, vol. 46, issue 01, pp165-190

[6] Joe, H. (1996): "Families of m-variate distributions with given margins and m(m/1)=2 bivariate dependence parameters. " Lecture Notes-Monograph Series 28, pp120–141.

References(2/2)

[7] Lee RD, Carter LR (1992) :"Modeling and Forecasting U.S. Mortality.", Journal of the American Statistical Association, 87(419), pp659-671.

[8] Li, N., and R. Lee (2005) :"Coherent mortality forecasts for a group of population: An extension to the classical Lee-Carter Approach" Demography 42: pp 575-594.

[9] Li, J. S.-H. and M. R. Hardy (2011) : "Measuring basis risk in longevity hedges" North American Actuarial Journal 15(2): pp 177-200.

[10] Lutz F. Gruber (n.d.) :"Review of Dependence Modeling with Regular Vine Copulas and Current Methods for Inference and Model Selection" (https://www.statistics.

ma.tum.de/fileadmin/w00bdb/www/LG/vine-copula-selection-inference-review.pdf)

[11] Rui Zhou (2019) : "Modelling Mortality Dependence With Regime-Switching Copulas", ASTIN Bulletin: The Journal of the International Actuarial Association, 2019, vol. 49, issue 02, pp373-407

[12] Sklar, A. (1959) : "Fonctions de Répartition à n Dimensions et Leurs Marges" Publications de l'Institut Statistique de l'Université de Paris, 8, pp229-231.

[13] Wenjun Zhu, Ken Seng Tan and Chou - Wen Wang (2017) :"Modeling Multi-Country Longevity Risk with Mortality Dependence: A Lévy Subordinated Hierarchical Archimedean Copulas (LSHAC) Approach", Journal of Risk and Insurance, 84, (S1), pp 477-493

Appendix



To examine the robustness, We apply same estimation procedure for the following three patterns.

Case 1 uses female data of the same age and period

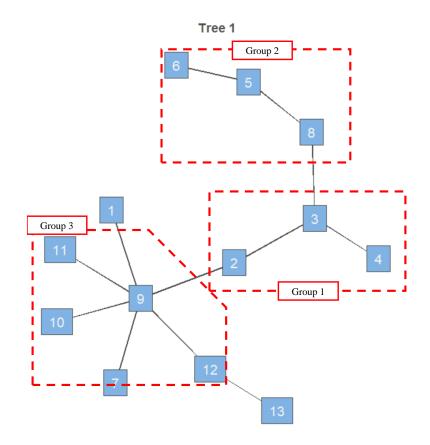
≻Case 2 uses the post-war period (1955-2014)

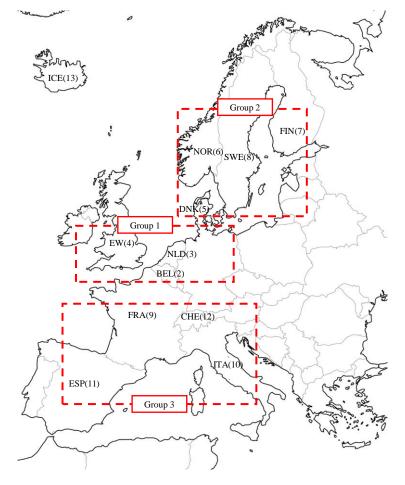
➤Case 3 uses working-age population (15-64)

	Sex	Year	Age
Case 1	Female	1921-2014	65-90
Case 2	Male	1955-2014	65-90
Case 3	Male	1921-2014	15-64



Case 1: Using female data of the same age and period



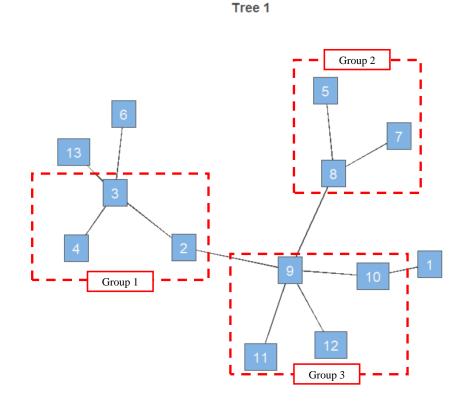


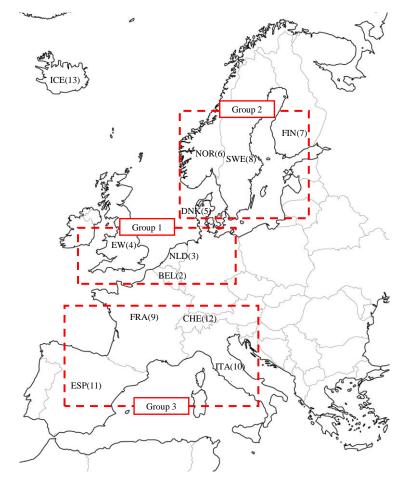
1 <-> AUS, 2 < -> BEL, 3 < -> NLD, 4 <-> EW, 5 < -> DNK, 6 < -> NOR, 7 < -> FIN, 8 <-> SWE, 9 < -> FRA, 10 < -> ITA, 11 < -> ESP, 12 < -> CHE, 13 < -> ICE

Aside form Finland(7), the dependence structure resembles the actual geographical location.



Case 2: Using the post-war period (1955-2014)



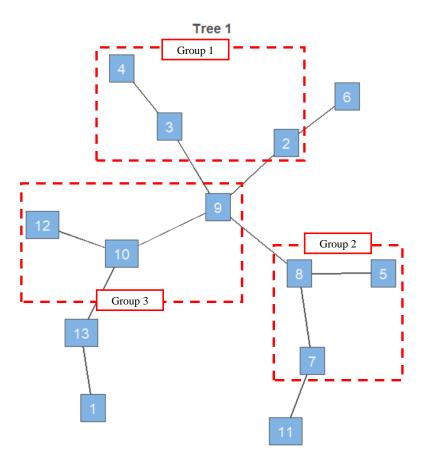


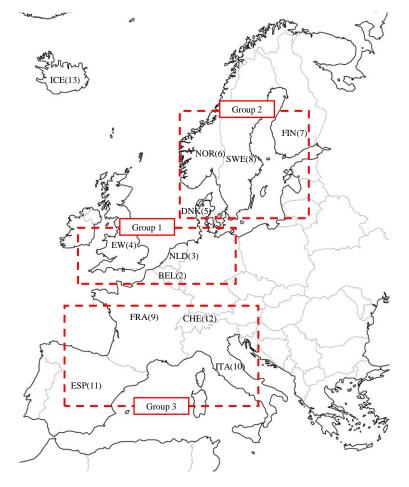
$$\begin{split} 1 <> \text{AUS}, \ 2 < \ -> \text{BEL}, \ 3 < \ -> \text{NLD}, \ 4 < \ -> \text{EW}, \ 5 < \ -> \text{DNK}, \ 6 < \ -> \text{NOR}, \ 7 < \ -> \text{FIN}, \\ 8 <-> \text{SWE}, \ 9 < \ -> \text{FRA}, \ 10 < \ -> \text{ITA}, \ 11 < \ -> \text{ESP}, \ 12 < \ -> \text{CHE}, \ 13 < \ -> \text{ICE} \end{split}$$

Aside form Norway(6), the dependence structure resembles the actual geographical location.



Case 3: Using working-age population (15-64)





1 <-> AUS, 2 < -> BEL, 3 < -> NLD, 4 < -> EW, 5 < -> DNK, 6 < -> NOR, 7 < -> FIN, 8 <-> SWE, 9 < -> FRA, 10 < -> ITA, 11 < -> ESP, 12 < -> CHE, 13 < -> ICE

Aside form Norway (6) and Spain (11), the dependence structure resembles the actual geographical location.